

# Modeling non-Gaussian spatial data

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## Non-Gaussian spatial data

- Often data sets preclude Gaussian modeling:  $y(\mathbf{s})$  may not even be continuous
- Examples:
  - Binary: presence or absence of a species at location  $\mathbf{s}$ .
  - Count: abundance of a species at location  $\mathbf{s}$ .
  - Categorical: counts of trees by size class at location  $\mathbf{s}$ .
- Replace Gaussian likelihood by exponential family member (Diggle, Tawn, and Moyeed (1998)).

# Hierarchical Bayesian approach

- **First stage:**  $y(\mathbf{s}_i)$  are conditionally independent given  $\beta$  and  $w(\mathbf{s}_i)$ . Here we use a canonical link function, say  $g(E[y(\mathbf{s}_i)]) = \eta(\mathbf{s}_i) = \mathbf{x}(\mathbf{s}_i)^\top \beta + w(\mathbf{s}_i)$ .

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- **Third stage:** Priors and hyperpriors.

# MCMC sampling for spatial GLMMs

- Additional GLMM flexibility comes at a computational cost: lose conjugacy of  $\beta, \mathbf{w}$
- Requires more Metropolis steps. Particularly costly for  $\mathbf{w}$
- Practical consequence: slower, less efficient algorithms
- Prediction and interpolation proceed as with the Gaussian case

## Binomial Spatial GLMMs

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- Both yield closed form full conditional distributions for all parameters except  $\phi$ .

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- Define  $\kappa(\mathbf{s}_i) = y(\mathbf{s}_i) - N(\mathbf{s}_i)/2$
- Resulting Gibbs sampler is remarkably similar to that of a Gaussian model with response  $y(\mathbf{s}_i)^* = \kappa(\mathbf{s}_i)/\omega(\mathbf{s}_i)$  and heteroskedastic variances  $\tau^2(\mathbf{s}_i) = 1/\omega(\mathbf{s}_i)$ .

## Pólya-Gamma data augmentation

- Suppose  $y(\mathbf{s}_i) \sim \text{Bernoulli}(\psi(\mathbf{s}_i))$ .

$$\begin{aligned}\psi(\mathbf{s}_i)^{y(\mathbf{s}_i)}(1 - \psi(\mathbf{s}_i))^{1-y(\mathbf{s}_i)} &= \frac{\exp(\mathbf{x}(\mathbf{s}_i)^\top \boldsymbol{\beta} + w(\mathbf{s}_i))^{y(\mathbf{s}_i)}}{1 + \exp(\mathbf{x}(\mathbf{s}_i)^\top \boldsymbol{\beta} + w(\mathbf{s}_i))} \\ &= \exp(\kappa(\mathbf{s}_i)(\mathbf{x}(\mathbf{s}_i)^\top \boldsymbol{\beta} + w(\mathbf{s}_i))) \times \\ &\quad \int \exp\left(-\frac{\omega(\mathbf{s}_i)}{2}(\mathbf{x}(\mathbf{s}_i)^\top \boldsymbol{\beta} + w(\mathbf{s}_i))\right)^2 \times \\ &\quad p(\omega(\mathbf{s}_i) \mid 1, 0) d\omega(\mathbf{s}_i),\end{aligned}$$

- $p(\omega(\mathbf{s}_i) \mid 1, 0)$  is the Pólya-Gamma PDF with parameters 1 and 0
- With Gaussian priors on  $\boldsymbol{\beta}$  and IG prior on  $\sigma^2$ , full conditionals for  $\boldsymbol{\beta}$ ,  $\sigma^2$ , and  $\mathbf{w}$  are available in closed form.  $\phi$  updated with MH.
- See Polson, Scott, Windle (2013) JASA**



## Example: species distribution modeling

- Objective: predict the distribution of Loggerhead Shrike across the US

$$y(\mathbf{s}_i) \sim \text{Bernoulli}(\psi(\mathbf{s}_i))$$

$$\text{logit}(\psi(\mathbf{s}_i)) = \mathbf{x}(\mathbf{s}_i)^\top \boldsymbol{\beta} + w(\mathbf{s}_i)$$

$$\mathbf{w} \sim N(\mathbf{0}, \sigma^2 \mathbf{R}(\phi))$$

$$\boldsymbol{\beta} \sim N(\boldsymbol{\mu}_\beta, \boldsymbol{\Sigma}_\beta)$$

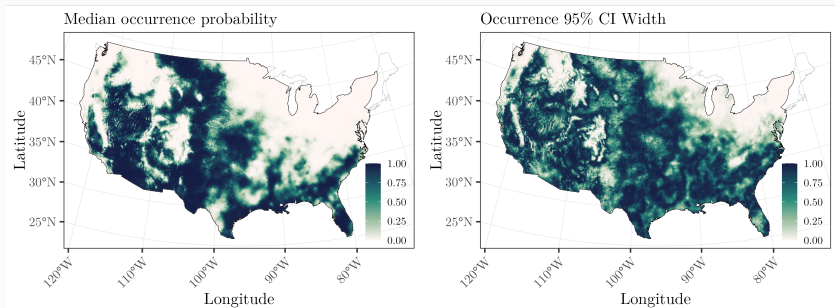
$$\sigma^2 \sim \text{IG}(a_\sigma, b_\sigma)$$

$$\phi \sim \text{Uniform}(l, u)$$

$$\omega(\mathbf{s}_i) \sim \text{PG}(1, 0)$$

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Posterior predictive inference proceeds as with the Gaussian case



## Some practical considerations

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- Be careful with non-identity link functions when thinking about priors.
- Pólya-Gamma data augmentation also applicable for Negative Binomial count data, but slow for large counts and can be unstable.

- spBayes
  - Univariate and multivariate, full GPs or predictive processes
  - Gaussian, Binomial (no Pólya-Gamma data augmentation), Poisson
- spNNGP
  - Univariate, NNGPs
  - Gaussian, Binomial
- spOccupancy
  - Univariate and multivariate, focus on modeling wildlife distributions, full GPs or NNGPs
  - Bernoulli
- spAbundance  
(<https://github.com/doserjef/spAbundance>)
  - Univariate and multivariate, focus on modeling wildlife/plant abundance, NNGPs
  - Gaussian, Poisson, Negative Binomial